SGD with Coordinate Sampling: Theory and Practice

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Introduction

Underlying optimization problem

Let $f:\mathbb{R}^p\to\mathbb{R}$ be a general objective function.

• Goal: Solve

$$\min_{\theta \in \mathbb{R}^p} \left\{ f(\theta) = \mathbb{E}_{\xi}[f(\theta, \xi)] \right\}$$

• **Constraints:** ∇f is hard to compute (large-scale problems) or even intractable (black-box) !

• Central question: Fast and Efficient procedures

Empirical Risk Minimization. data $z_1, \ldots, z_n \subset \mathcal{Z}$ and loss function $\ell : \mathbb{R}^p \times \mathcal{Z} \to \mathbb{R}$,

$$\forall \theta \in \mathbb{R}^p, \quad f(\theta) = \frac{1}{n} \sum_{i=1}^n \ell(\theta, z_i)$$

The true gradient, $n^{-1}\sum_{i=1}^n \nabla \ell(\theta,z_i)$ requires n evaluations

Introduction

Noisy gradients

• Zeroth-Order (biased):

$$g(\theta) = \sum_{k=1}^{p} h^{-1} (f(\theta + he_k) - f(\theta)) e_k \underset{h \to 0}{\approx} \nabla f(\theta)$$

• First-Order (unbiased):

$$\boldsymbol{g}_{t+1} := \nabla_{\boldsymbol{\theta}} \ell(\boldsymbol{\theta}_t, \boldsymbol{z}_{\xi_{t+1}})$$

where $\xi_{t+1} \sim \mathcal{U}(\llbracket 1, n \rrbracket)$ is uniformly distributed.

Stochastic Gradient Descent (Robbins and Monro, 1951)

- Start (t = 0) from random point $\theta_0 \in \mathbb{R}^d$.
- Evaluate noisy gradient g_{t+1}
- Update iterate $\theta_{t+1} = \theta_t \gamma_{t+1} g_{t+1}$.

Introduction

• (SCGD): Stochastic Coordinate Gradient Descent

$$\begin{cases} \theta_{t+1}^{(k)} = \theta_t^{(k)} & \text{if } k \neq \zeta_{t+1} \\ \theta_{t+1}^{(k)} = \theta_t^{(k)} - \gamma_{t+1} g_{t+1}^{(k)} & \text{if } k = \zeta_{t+1} \end{cases}$$

 ζ_{t+1} is a random variable valued in $\llbracket 1, p \rrbracket$.

- Reduction of the computing cost
- Covers all approaches that uses a gradient estimate g_{t+1}
- 2 sources of randomness:
 (i) noisy gradient g_{t+1}
 (ii) noisy coordinate ζ_{t+1}

• (SCGD): Stochastic Coordinate Gradient Descent

$$\begin{cases} \theta_{t+1}^{(k)} = \theta_t^{(k)} & \text{if } k \neq \zeta_{t+1} \\ \theta_{t+1}^{(k)} = \theta_t^{(k)} - \gamma_{t+1} g_{t+1}^{(k)} & \text{if } k = \zeta_{t+1} \end{cases}$$

- How to update the selecting policy ζ_{t+1}?
 → We develop an algorithm MUSKETEER to leverage the data structure and move along relevant directions.
- What condition on ζ_{t+1} for convergence ?
 → We analyze the properties of SCGD algorithms (convergence of the iterates, convergence of the policy, non-asymptotic bound)

- CD using f or true gradient ∇f (Loshchilov et al., 2011; Richtárik and Takáč, 2013; Glasmachers and Dogan, 2013; Qu and Richtárik, 2016; Allen-Zhu et al., 2016; Namkoong et al., 2017)
- Most related idea: **Gauss-Southwell rule** to select the largest gradient coordinate to move the iterate (Nutini et al., 2015) \rightarrow Here: we have stochastic g_{t+1} and ζ_{t+1} .
- Sparsification methods (Alistarh et al., 2017; Wangni et al., 2018), unbiased importance sampling estimate of the gradient → Here: no reweighting (biased) (conditioned gradient)

General framework and notation

• Only one coordinate ζ_{t+1} is selected:

$$(SCGD) \quad \theta_{t+1} = \theta_t - \gamma_{t+1} D(\zeta_{t+1}) g_{t+1}$$

with $D(k)=e_ke_k^T=Diag(0,\ldots,0,1,0,\ldots,0).$

• The distribution of ζ_{t+1} , is the **coordinate sampling policy** and is given by the probability weights vector $d_t = (d_t^{(1)}, \ldots, d_t^{(p)})$

$$d_t^{(k)} = \mathbb{P}(\zeta_{t+1} = k | \mathcal{F}_t), \quad k \in [\![1, p]\!].$$

 Not the same mean field as in usual SGD. Under conditional independence between g_{t+1} and ζ_{t+1}:

$$\mathbb{E}[D(\zeta_{t+1})g_{t+1}|\mathcal{F}_t] = \operatorname{diag}(d_t)g(\theta_t)$$

General update rule

$$\theta_{t+1} = \theta_t - \gamma_{t+1} h(\theta_t, \omega_{t+1})$$

where h is a gradient generator and $(\omega_t)_{t\geq 1}$ is a sequence of random variables

- (SGD) $h(\theta, \omega_{t+1}) = g_{t+1}$
- (SCGD) $h(\theta, \omega_{t+1}) = D(\zeta_{t+1})g_{t+1}$
- (Unbiased with importance weights as in (Wangni et al., 2018)) $h(\theta, \omega_{t+1}) = D_t^{-1} D(\zeta_{t+1}) g_{t+1}$

MUSKETEER

MUltivariate Stochastic Knowledge Extraction Through Exploration Exploitation Reinforcement



Illustration/Motivation



MUSKETEER may be seen as an adaptive bandit problem with

'arms = coordinates'

Alternate between 2 phases

• Exploration phase (one for all)

- \rightarrow fixed $d_t\text{,}$ draw random coordinate and move along selected direction
- \rightarrow cumulative gains for the visited coordinates
- Exploitation phase (all for one)
- \rightarrow share knowledge of the cumulative gains
- ightarrow update the coordinate sampling probability vector d_t

1) Pick a coordinate

Generate $\zeta_{t+1} \sim d_t$ and the coordinate gradient g_{t+1} 2) Update the iterate

$$\theta_{t+1}^{(\zeta_{t+1})} = \theta_t^{(\zeta_{t+1})} - \gamma_{t+1} g_{t+1}^{(\zeta_{t+1})}$$

3) Update cumulative gains

$$G_{t+1}^{(\zeta_{t+1})} = G_t^{(\zeta_{t+1})} + \frac{g_{t+1}^{(\zeta_{t+1})}}{d_t} / d_t^{(\zeta_{t+1})}$$

- \rightarrow (Variants with square) $|g_{t+1}^{(\zeta)}|$ or $g_{t+1}^{(\zeta)2}$
- \rightarrow Might be done T times with d_t fixed (before moving to the exploitation)

- This phase is to update the policy value of d_t
- EXP3 algorithm (Auer et al., 2002) to update the probability weights through a mixture. Given $\eta > 0$ and $\lambda \in [0, 1]$, we have for all $k \in [\![1, p]\!]$,

$$d_{t+1}^{(k)} = (1-\lambda) \frac{\exp(\eta |G_{t+1}^{(k)}|/t)}{\sum_{j=1}^{d} \exp(\eta |G_{t+1}^{(j)}|/t)} + \lambda \frac{1}{p}$$

• The mixture with $\lambda > 0$ ensure to always give a chance to everyone

Algorithm input: (d_0, θ_0) , sequence $(\gamma_t)_{t \ge 1}$ and parameter (η, λ)

- 1: for t = 0, 1, 2, ... do
- 2: Set $d = d_t$ and sample coordinate $\zeta \sim d$ and gradient g
- 3: Update iterate: $\theta_{t+1}^{(\zeta)} = \theta_t^{(\zeta)} \gamma_{t+1}g^{(\zeta)}$
- 4: Update gain: $G_{t+1}^{(\zeta)} = G_t^{(\zeta)} + g^{(\zeta)}/d^{(\zeta)}$
- 5: Whenever $t = 0 \pmod{T}$: update weights d_{t+1} with

$$d_{t+1}^{(k)} = (1 - \lambda) \frac{\exp(\eta |G_t^{(k)}|/t)}{\sum_{j=1}^d \exp(\eta |G_t^{(j)}|/t)} + \lambda \frac{1}{p}$$

6: end for

Numerical Experiments

• We apply ERM to regularized regression and classification problems.

• Given a data matrix $X = (x_{i,j}) \in \mathbb{R}^{n \times p}$ with labels $y \in \mathbb{R}^n$ and a regularization parameter $\mu > 0$, the *Ridge regression* is

$$\min_{\theta \in \mathbb{R}^p} f(\theta) = \frac{1}{2n} \sum_{i=1}^n (y_i - \sum_{j=1}^p x_{i,j} \theta_j)^2 + \frac{\mu}{2} \|\theta\|_2^2$$

and the ℓ_2 -regularized logistic regression is defined by

$$\min_{\theta \in \mathbb{R}^p} f(\theta) = \frac{1}{n} \sum_{i=1}^n \log(1 + \exp(-y_i \sum_{j=1}^p x_{i,j} \theta_j)) + \mu \|\theta\|_2^2$$

where μ is set to the classical value $\mu=1/n$

Special covariance structure $X[:,k] \sim \mathcal{N}(0, \sigma_k^2 I_n)$ with $\sigma_k^2 = k^{-\alpha}$ for $k \in [\![1,p]\!]$

Setting $\gamma_t = 1/t$, n = 10,000, p = 250, $T = \lfloor \sqrt{p} \rfloor = 15$

ZO Ridge Regression ($\alpha = 5$ and $\alpha = 10$)



ZO Logistic Regression ($\alpha = 2$ and $\alpha = 5$)



Numerical Experiments

• MNIST and Fashion-MNIST (ZO) (p = 55, 050 and T = 234)



Back ground

Stochastic Optimization

$$\min_{\theta \in \mathbb{R}^p} \left\{ f(\theta) = \mathbb{E}_{\xi}[f(\theta, \xi)] \right\}$$

Gradients might be biased

There exists constant $c \ge 0$ such that

$$\forall h > 0, \theta \in \mathbb{R}^p, \quad \|\mathbb{E}_{\xi}[g_h(\theta, \xi)] - \nabla f(\theta)\| \le ch.$$

- $h \ge 0$ is a parameter controlling the bias
- c = 0 recovers 1st-order gradient estimates
- Allows to cover general zeroth-order estimates

ZO gradient estimates

Example 1 (smoothing).

(Nesterov and Spokoiny, 2017). The smoothed gradient estimate is

$$\forall \theta \in \mathbb{R}^p, \boldsymbol{g}_h(\theta, \xi) = h^{-1} [f(\theta + hU, \xi) - f(\theta, \xi)] U$$

where $U \sim \mathcal{N}(0, I)$ (Alternative version with $U \sim Unif(\mathbb{S})$)

Example 2 (finite differences).

The finite differences gradient estimate is given by

$$\forall \theta \in \mathbb{R}^p, g_h(\theta, \xi) = \sum_{k=1}^p g_h(\theta, \xi)^{(k)} e_k$$

where for all $k = 1, \ldots, p$ the coordinates are

$$g_h(\theta,\xi)^{(k)} = h^{-1}[f(\theta + he_k,\xi) - f(\theta,\xi)]$$

General form

There exists probability measure ν satisfying $\int_{\mathbb{R}^p} x x^\top \nu(\mathrm{d}x) = I_p$,

$$\forall h > 0, \theta \in \mathbb{R}^p, \quad \mathbb{E}_{\xi}[g_h(\theta, \xi)] = \int_{\mathbb{R}^p} x \left\{ \frac{f(\theta + hx) - f(\theta)}{h} \right\} \nu(\mathrm{d}x).$$

Lemma

Under the previous assumption (if f is L-smooth) the biased gradient assumption is satisfies with

$$c = (L/2) \sqrt{\int_{\mathbb{R}^p} \|x\|_2^6 \nu(\mathrm{d}x)}$$

- \bullet smoothing gradient is recovered when ν is the Gaussian measure
- Take $\nu = \sum_{k=1}^p \delta_{e_k}/p$ covers the finite differences estimate
- (MUSKETEER) Use a measure ν that evolves through time and put different weights on the different directions !

Assumption

Growth condition

There exist $0 \leq \mathcal{L}, \sigma^2 < \infty$

 $\forall h > 0, \theta \in \mathbb{R}^p \quad \mathbb{E}\left[\| \boldsymbol{g}_h(\theta, \xi) \|_{\infty}^2 \right] \le 2\mathcal{L}\left(f(\theta) - f^{\star} \right) + \sigma^2.$

Smoothness and lower bound

f is $L\mbox{-smooth}$ and lower bounded by f^\star

Two algorithms

Gradient generator $g_t = g_{h_{t+1}}(\theta_t, \xi_{t+1})$

$$(SGD) \quad \theta_{t+1} = \theta_t - \gamma_{t+1}g_t$$
$$(SCGD) \quad \theta_{t+1} = \theta_t - \gamma_{t+1}D(\zeta_{t+1})g_t$$

Robbins-Monro $\sum_{t\geq 1}\gamma_t=+\infty \quad \text{and} \quad \sum_{t\geq 1}\gamma_t^2<+\infty$ small bias $h_t^2=O(\gamma_t)$

Theorem (Almost sure convergence of (biased) SGD) Under previous assumptions, $\nabla f(\theta_t) \rightarrow 0$ a.s. when $t \rightarrow \infty$.

Theorem (Almost sure convergence of particular SCGD)

Under previous assumptions

- (i) (max gradient) if $\zeta_{t+1} = \arg \max_{k=1,...,p} |\partial_k f(\theta_t)|$ then $\nabla f(\theta_t) \to 0$ almost surely as $t \to +\infty$.
- (ii) (gradient weights) if $D_t \propto (|\nabla_k f(\theta_t)|^q)_{1 \le k \le p}$ with q > 0 then $\nabla f(\theta_t) \to 0$ almost surely as $t \to +\infty$.

• When f coercive and unique solution $\{\theta : \nabla f(\theta) = 0\} = \{\theta^*\}$ then almost sure convergence towards minimizer $\theta_t \to \theta^*$.

Theorem (Almost sure convergence general SCGD)

Under previous assumptions, if $\beta_{t+1} = \min_{1 \le k \le p} d_t^{(k)}$ is away from 0 then $\nabla f(\theta_t) \to 0$ almost surely as $t \to +\infty$.

Theorem (Almost sure convergence)

The sequence of iterates $(\theta_t)_{t\geq 0}$ obtained by the MUSKETEER satisfies $\nabla f(\theta_t) \to 0$ almost surely as $t \to +\infty$.

Theorem (Weak convergence)

The MUSKETEER's coordinate policy $(d_t)_{t\in\mathbb{N}}$ converges weakly to the uniform distribution

Theorem (Non-asymptotic bounds, (Moulines and Bach, 2011))

Let $(\theta_t)_{t\in\mathbb{N}}$ obtained by MUSKETEER with $\gamma_t = \gamma t^{-\alpha}$ then

$$\mathbb{E}\left[f(\theta_t) - f^\star\right] = O(1/t), \quad (\alpha = 1)$$

Contributions

• (**Theory**) Almost-sure convergence SCGD towards stationary points, non-asymptotic bounds on the optimality gap $\mathbb{E}[f(\theta_t) - f^*]$.

• Conditions are relatively weak as f is only L-smooth (classical in non-convex problems) and the stochastic gradients are possibly biased with unbounded variance.

(Practice) New algorithm, called MUSKETEER: in the image of the motto 'all for one and one for all', this procedure belongs to the SCGD framework with a particular design for the *coordinate sampling policy*.
MUSKETEER compares the value of all past gradient estimates g_t to select a descent direction (*all for one*) and then moves the current iterate according to the chosen direction (*one for all*).

Future work

Study the asymptotic behavior of other adaptive sampling strategies

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