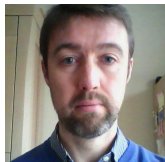


A Quadrature Rule combining Control Variates and Adaptive Importance Sampling

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Background, Goal and Contributions

- Sequential simulation has emerged as a leading approach to compute multidimensional integrals where the **target density** may be known only up to a proportionality constant (e.g. Bayesian inference)
- While the design of algorithms with adaptive **policies** has been of major interest recently, only few studies have focused on **control variates** to reduce the variance.

GOAL: numerically calculate an integral using **importance sampling** and reduce the variance by including **control variates**.

Contributions:

- (1) A simple weighted least squares approach is proposed to improve the procedure of sequential algorithms with control variates.
- (2) The proposed approach significantly improves the accuracy of the initial algorithm, both theoretically and in practice.
- (3) It takes the form of a **quadrature rule** with adapted quadrature weights that **do not depend on the integrand** and reflect the information brought in by the control variates.
- (4) **Non-asymptotic bound** on the probabilistic error of the procedure.

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Preliminaries on Monte Carlo Integration

GOAL: Given an **integrand** $g : \mathbb{R}^d \rightarrow \mathbb{R}$ and a target **density function** f , the goal is to compute the integral

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Can we sample from target distribution f ?

- **YES**, then use naive Monte Carlo estimate (later on control variates)

$$I_n^{(\text{mc})}(g) = \frac{1}{n} \sum_{i=1}^n g(X_i), \quad X_1, \dots, X_n \sim f$$

- **NO**, then use **importance sampling** with sampling policy q

$$I_{\text{norm}}^{(\text{is})}(g) = \frac{\sum_{i=1}^n w_i g(X_i)}{\sum_{i=1}^n w_i}, \quad X_1, \dots, X_n \sim q,$$

where the sequence $(w_i)_{i=1, \dots, n}$ of **importance weights** is defined by

$$w_i = f(X_i)/q(X_i).$$

Generalization: Adaptive Importance Sampling (AIS)

GOAL:

$$\alpha = \mathbb{E}_f[g] = \int_{\mathbb{R}^d} g(x)f(x) dx$$

- Use a **sampling policy** $(q_t)_{t \geq 0} =$ a sequence of densities which evolves adaptively depending on previous outcomes with $q_t \rightarrow f$ when $t \rightarrow \infty$. and an **allocation policy** $(n_t)_{t \geq 0}$.
- At time t , draw n_t **particles** $X_{t,1}, \dots, X_{t,n_t} \sim q_{t-1}$ with **importance weights** $w_{t,i} = f(X_{t,i})/q_{t-1}(X_{t,i})$.
- The normalized AIS estimate [DP21] of α is given by

$$I_{\text{norm}}^{(\text{ais})}(g) = \frac{\sum_{t=1}^T \sum_{i=1}^{n_t} w_{t,i} g(X_{t,i})}{\sum_{t=1}^T \sum_{i=1}^{n_t} w_{t,i}}.$$

Control Variates: variance reduction with samples from f

GOAL:

$$\alpha = \mathbb{E}_f[g] = \int_{\mathbb{R}^d} g(x)f(x) dx$$

• **Control variates** are functions $h_1, \dots, h_m \in L_2(f)$ with known integrals. Let $h = (h_1, \dots, h_m)^\top$, assume that $\mathbb{E}_f[h_j] = 0$ for all $j = 1, \dots, m$. (Stein control variates)

• For any $\beta \in \mathbb{R}^m$, we have $\mathbb{E}_f[g - \beta^\top h] = \mathbb{E}_f[g]$ leading to the CV estimate of α , parameterized by β

$$I_n^{(\text{cv})}(g, \beta) = \frac{1}{n} \sum_{i=1}^n [g(X_i) - \beta^\top h(X_i)], \quad X_1, \dots, X_n \sim f.$$

• What optimal choice for β^* ? Look at variance and define

$$\beta^* = \arg \min_{\beta \in \mathbb{R}^m} \mathbb{E}_f [(g - \mathbb{E}_f[g] - \beta^\top h)^2]$$

Control Variates and Least-Squares

- Provided matrix $G = \mathbb{E}_f[hh^\top]$ is invertible, there is a unique $\beta^* \in \mathbb{R}^m$ for which the variance of $I_n^{(cv)}(g)$ is minimal: $\beta^* = (\mathbb{E}_f[hh^\top])^{-1} \mathbb{E}_f[hg]$.
- Casting the problem in an **Ordinary Least Squares** framework leads to the control variate estimate

$$I_n^{(cv)}(g) = I_n^{(cv)}(g, \hat{\beta}_n^{(cv)}) = \hat{a}_n^{(cv)} \quad \text{where } X_1, \dots, X_n \sim f,$$

$$(\hat{a}_n^{(cv)}, \hat{\beta}_n^{(cv)}) \in \arg \min_{(a,b) \in \mathbb{R} \times \mathbb{R}^m} \frac{1}{n} \sum_{i=1}^n \{g(X_i) - a - b^\top h(X_i)\}^2$$

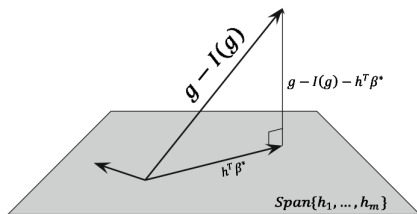


Figure: L^2 projection of g onto space of control variates $\text{Span}\{h_1, \dots, h_m\}$.

Adaptive Importance Sampling with Control Variates

- AISCV estimate is the first coordinate of the solution to the **Weighted Least Squares** problem

$$(\hat{\alpha}_n, \hat{\beta}_n) = \arg \min_{a \in \mathbb{R}, b \in \mathbb{R}^m} \sum_{i=1}^n w_i \left[g(X_i) - a - b^\top h(X_i) \right]^2, \quad w_i = f(X_i) / q_{i-1}(X_i).$$

- (a) (Exact integration) whenever g is of the form $\alpha + \beta^\top h$ for some $\alpha \in \mathbb{R}$ and $\beta \in \mathbb{R}^m$, the **error is zero**, i.e., $\hat{\alpha}_n = \alpha = \int g f \, d\lambda$.
- (b) (Quadrature Rule) the estimate takes the form of a quadrature rule $\hat{\alpha}_n = \sum_{i=1}^n v_{n,i} g(X_i)$, for **quadrature weights** $v_{n,i}$ that **do not depend on the function** g and that can be computed by a single weighted least squares procedure.
- (c) (Bayesian) it can be computed even when f is **known only up to a multiplicative constant**.
- (d) (post-hoc scheme) CV can be brought into play in a **post-hoc** scheme, after generation of the particles and importance weights, and **this for any AIS algorithm**

Concentration inequality for the AISCV estimate

Theorem

Under assumptions, for any $\delta \in (0, 1)$ and for all $n \geq C_1 c^2 B \log(10m/\delta)$, we have, with probability at least $1 - \delta$, that

$$\left| I_{\text{norm}}^{(\text{aiscv})}(g) - \int_{\mathbb{R}^d} g(x) f(x) dx \right| \leq C_2 \tau \sqrt{\frac{\log(10/\delta)}{n}} + C_3 c B \tau \frac{\log(10m/\delta)}{n},$$

where C_1, C_2, C_3 are some constants and $B = \sup_{x: f(x) > 0} \|\tilde{h}(x)\|_2^2$,
 $\tilde{h} = G^{-1/2} h$.

Control Variates in Practice and Bayesian Inference

- **Stein control variates** [OGC17] are built with operator \mathcal{L} on functions $\varphi \in \mathcal{C}^2(\mathbb{R}^d, \mathbb{R})$ to have $\mathbb{E}_f[\mathcal{L}\varphi] = 0$.

$$(\mathcal{L}\varphi)(x) = \Delta_x \varphi(x) + \nabla_x \varphi(x)^\top \nabla_x \log f(x).$$

- $\nabla_x \log f(x)$ can either be directly computed (Bayesian regression) or with autodiff (Tensorflow and PyTorch).
- Given data \mathcal{D} and parameter of interest $\theta \in \mathbb{R}^d$, posterior integrals take the form $\int_{\mathbb{R}^d} g(\theta) p(\theta|\mathcal{D}) d\theta$, where $p(\theta|\mathcal{D}) \propto \ell(\mathcal{D}|\theta)\pi(\theta)$ is the posterior distribution, proportional to prior $\pi(\cdot)$ and a likelihood $\ell(\mathcal{D}|\cdot)$.

Synthetic examples: Gaussian Mixtures

Integrand and Target: $g(x) = x$, $f_{\Sigma}(x) = 0.5\Phi_{\Sigma}(x - \mu) + 0.5\Phi_{\Sigma}(x + \mu)$
where $\mu = (1, \dots, 1)^{\top} / 2\sqrt{d}$, $\Sigma = I_d/d$ and Φ_{Σ} is pdf $\mathcal{N}(0, \Sigma)$.

Sampling policy: Multivariate Student

Control variates: Stein method with $\varphi =$ polynomial with bounded degree

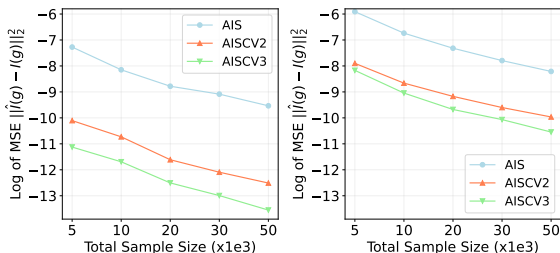


Figure: Gaussian mixture density: Logarithm of $\|\hat{I}(g) - I(g)\|_2^2$ for $g(x) = x$ with target isotropic f_{Σ} with $d = 4$ (left), $d = 8$ (right).

Bayesian Linear Regression on Real-world data

Data [DG19]: *housing* ($N = 506$; $d = 13$; $m \in \{12; 104\}$); *abalone* ($N = 4177$; $d = 8$; $m \in \{7; 44\}$).

Prior: $\pi(\theta) \sim \mathcal{N}(\mu_a, \Sigma_a)$, **Posterior:** $p(\theta|\mathcal{D}) \propto \ell(\mathcal{D}|\theta)\pi(\theta)$.

Integrand: $g(\theta) = \sum_{i=1}^d \theta_i^2$.

Control variates: Stein control variates with $\varphi_\alpha(\theta) = \theta_1^{\alpha_1} \cdots \theta_d^{\alpha_d}$, $\alpha_1 + \cdots + \alpha_d \leq Q$, $Q \in \{1; 2\}$.

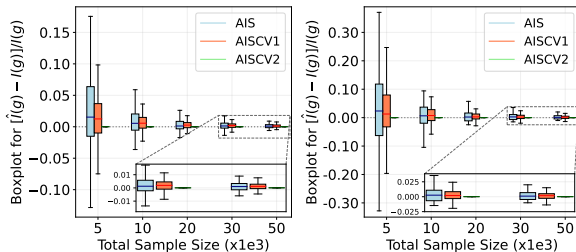


Figure: BLR: boxplots of $(\hat{I}(g) - I(g))/I(g)$ for $g(\theta) = \sum_{j=1}^d \theta_j^2$ with datasets Housing (left) and Abalone (right).

Conclusion and take-home message

- This paper provides a new method to incorporate **control variates** within standard **sequential algorithms**.
- The proposed approach significantly improves the accuracy of the initial algorithm, **both theoretically and in practice**.
- Control Variates can be brought into play in a ***post-hoc*** scheme, after generation of the particles and importance weights, and **this for any AIS algorithm**

Thank you and see you at the conference !

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