

Control Variates Selection for Monte-Carlo Integration

Rémi Leluc

Joint work with François Portier and Johan Segers



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- 1 Introduction
- 2 Mathematical Background
- 3 Monte-Carlo Control Variates
- 4 Non-Asymptotic Error Analysis
- 5 Numerical Experiments
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Introduction

- Solve deterministic problems by a stochastic approach

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- Simple, Flexible, Scalable

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- Many fields of applications that include: physical science, engineering, climate change, biology, applied statistics, artificial intelligence for games, finance and business.



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- Evaluate the function at nodes $f(X_1), \dots, f(X_n)$.
- Compute an approximation of $P(f)$ based on $((X_1, f(X_1)), \dots, (X_n, f(X_n)))$.

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- Let $X_1, \dots, X_n \stackrel{\text{i.i.d}}{\sim} P$ be an independent random sample from P on a probability space $(\Omega, \mathcal{F}, \mathbb{P})$. The naive Monte-Carlo estimator of $P(f)$ is given by the empirical mean

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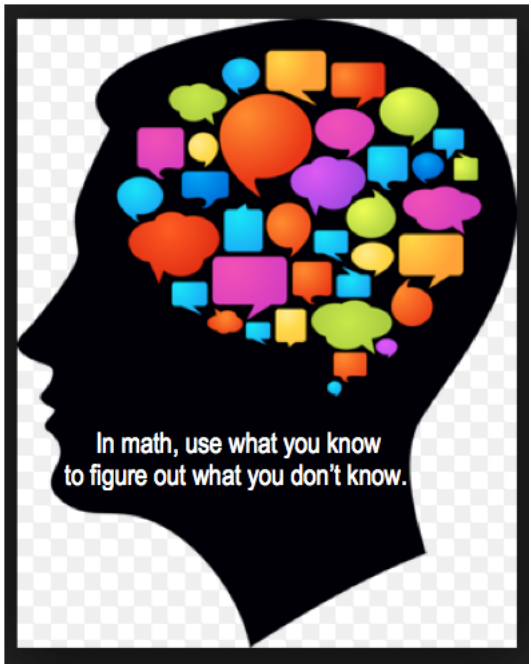
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- The Monte-Carlo estimator $P_n(f)$ of $P(f)$ is unbiased and has variance $\sigma^2(f)/n$, where $\sigma^2(f) = P[(f - P(f))^2]$. By the central limit theorem, we have the convergence in distribution

$$\sqrt{n}(P_n(f) - P(f)) \xrightarrow[n \rightarrow +\infty]{d} \mathcal{N}(0, \sigma^2(f))$$

- Antithetic variates
- Control variates

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- For $\beta = (\beta_1, \dots, \beta_m)^T \in \mathbb{R}^m$, $P(f - \beta^T h) = P(f)$, so $P_n(f - \beta^T h)$ is an unbiased estimator of $P(f)$.

Control Variates estimator

- Class of Monte-Carlo estimators

$$\hat{I}_n^{(\text{cv})}(f, \beta) = \frac{1}{n} \sum_{i=1}^n \left\{ f(X_i) - \beta^T h(X_i) \right\}, (\beta \in \mathbb{R}^m)$$

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- Minimize the variance to find optimal β

$$\beta^*(f) \in \arg \min_{\beta \in \mathbb{R}^m} P[(f - P(f) - \beta^T h)^2] = \arg \min_{\beta \in \mathbb{R}^m} \|(f - P(f)) - \beta^T h\|_{L^2}^2$$

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- If $\beta^*(f)$ would be known, the use of control variates would always reduce the variance of the basic Monte Carlo estimator.

Control Variates estimator

- The integral $P(f)$ thus appears as the intercept of a linear regression model with response f and explanatory variables h_1, \dots, h_m ,

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- (Hilbert projection) Normal equation: $P(hh^T)\beta^*(f) = P(hf)$
- Need to estimate the Gram matrix $P(hh^T)$ and $P(hf)$

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- $\mathbf{1}_n = (1, \dots, 1)^T \in \mathbb{R}^n$, $\|\cdot\|_2$ denotes the Euclidean norm and H is the random $n \times m$ matrix defined by $H = (h_j(X_i))_{\substack{i=1, \dots, n \\ j=1, \dots, m}}$.

$$\hat{\alpha}_n^{\text{ols}}(f) = P_n[f - \hat{\beta}_n^{\text{ols}}(f)^T h] = \sum_{i=1}^n w_{n,i} f(X_i)$$

Control Variates in the literature

- Stein method to build control functionals with non-parametric extension \hookrightarrow **Control functionals for Monte Carlo integration**, Oates, C. J., M. Girolami, and N. Chopin (2014)

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- Variance reduction via regularization \hookrightarrow **Regularised Zero-Variance Control Variates for High-Dimensional Variance Reduction**, South, L. F., C. J. Oates, A. Mira, and C. Drovandi (2018)

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Bet on sparsity
variable selection !

Lasso Monte-Carlo (LASSOMC)

- Adding ℓ_1 -penalization leads to

$$(\hat{\alpha}_n^{\text{lasso}}(f), \hat{\beta}_n^{\text{lasso}}(f)) = \arg \min_{(\alpha, \beta) \in \mathbb{R} \times \mathbb{R}^m} \frac{1}{2n} \|f^{(n)} - \alpha \mathbf{1}_n - H\beta\|_2^2 + \lambda \|\beta\|_1$$

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- Lasso takes advantage of *sparse* regression models. The *active set* associated to the coefficient vector $\beta \in \mathbb{R}^m$ is

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- The number of elements in $S^* = S(\beta^*(f))$, denoted by $\ell^* := |S^*|$, quantifies the level of sparsity associated to the regression model.

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- Lasso to select the active variables among a large number of control variates, then compute OLSMC using only the variables selected at the previous stage.
- Let $\hat{S} = \{k \in \{1, \dots, m\} : \hat{\beta}_{N,k}^{\text{lasso}}(f) > 0\}$ denote the estimated active set of control variates based on the subsample of size N . The LSLASSOMC estimate $\hat{\alpha}_n^{\text{llasso}}(f)$ of $P(f)$ is defined by

$$(\hat{\alpha}_n^{\text{llasso}}(f), \hat{\beta}_n^{\text{llasso}}(f)) \in \arg \min_{(\alpha, \beta) \in \mathbb{R} \times \mathbb{R}^{\hat{\ell}}} \|f^{(n)} - \alpha \mathbf{1}_n - H_{\hat{S}} \beta\|_2^2$$

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Assumptions: **sub-gaussian** residuals with factor τ , **linearly independent** and **bounded** control variates, appropriate λ

$$(U_h = \max_{j=1,\dots,m} \sup_{x \in \mathcal{X}} |h_j(x)|, \gamma = \lambda_{\min}(G), \zeta_h = U_h^2/\gamma)$$

Concentration inequalities (Leluc, Portier, Segers, 2019)

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- Methods in competition: OLSMC, LassoMC, LSLassoMC(X)

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d	k	Degree threshold				
		1	3	5	10	12
3	12	3	19	55	285	454
5	10	5	55	251	3 001	6 157
8	3	8	164	1 214	20 993	36 813

Table: Number of control variates by degrees

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n	N	$\lfloor 3\sqrt{n} \rfloor$	$\lfloor 12\sqrt{n} \rfloor$
2 000	700	134	536
5 000	1 000	212	848
10 000	2 000	300	1 200

Table: Parameters setting with range $(c_1\sqrt{n}, c_2\sqrt{n})$ of selected control variates.

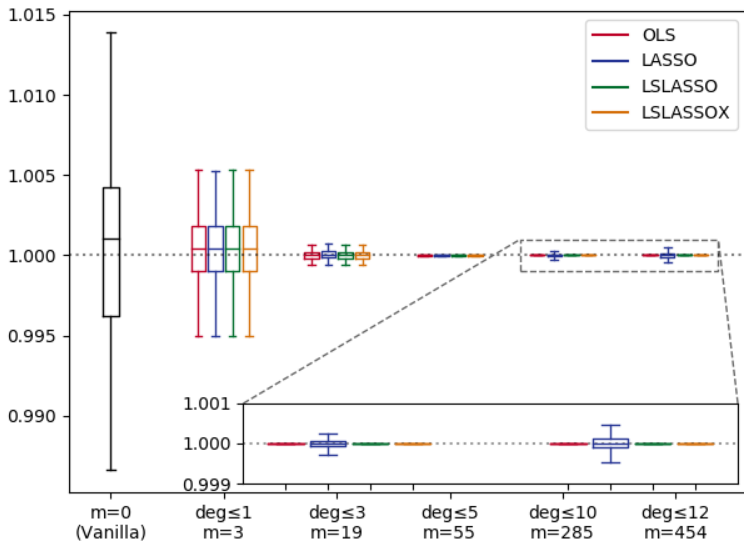


Figure: $\varphi, d = 3, n = 10\,000, N = 2\,000$.

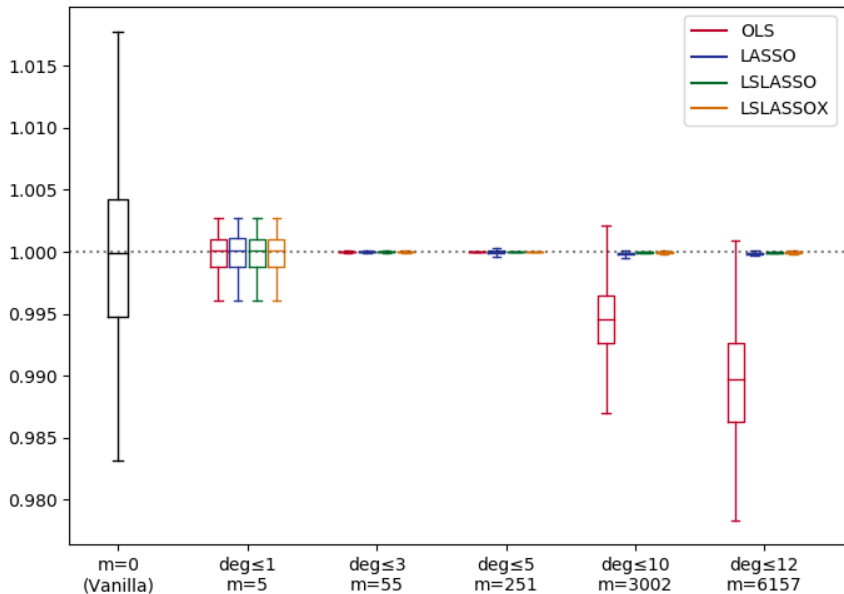


Figure: $g_3, d = 5, n = 2000, N = 700$.

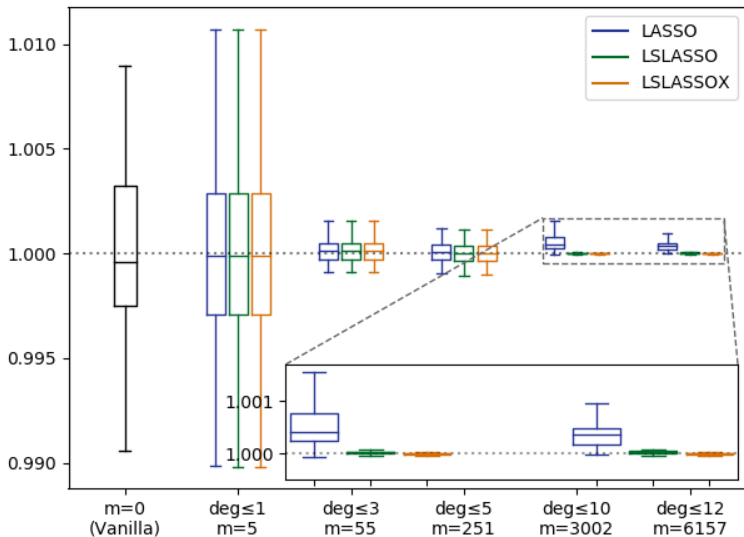


Figure: $f_1, d = 5, n = 5000, N = 1000$

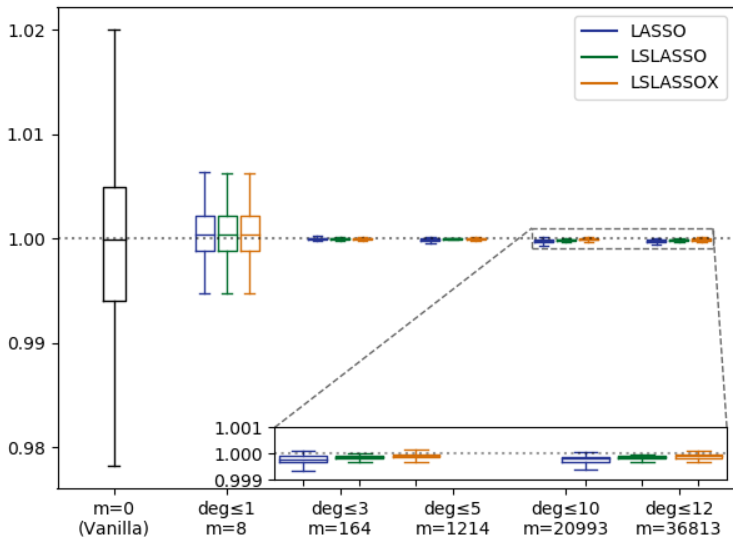


Figure: $g_4, d = 8, n = 2000, N = 700$.

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Conclusion

- The particular variance reduction technique of control variates offers many advantages as it relies on a simple and intuitive paradigm.

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- Regularizing the ordinary least squares estimator by preselecting appropriate control variates via the Lasso turns out to increase the accuracy without additional computational cost.

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- Regularizing the ordinary least squares estimator by preselecting appropriate control variates via the Lasso turns out to increase the accuracy without additional computational cost.
- The proposed numerical method performs better than any other state of the art method.

Questions and Answers



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- The control variates $h_1, \dots, h_m \in L^2(P)$ are **linearly independent**. As a consequence, the **Gram matrix** $G := P(hh^T) \in \mathbb{R}^m$ is **positive definite** and its smallest eigenvalue $\gamma := \lambda_{\min}(G)$ is positive.
- The residual function $\epsilon = f - P(f) - \beta^*(f)^T h$ satisfies $\epsilon \in \mathcal{G}(\tau^2)$ for some $\tau > 0$, that is, $\int_{\mathcal{X}} \exp(\lambda x) \epsilon(x) P(dx) \leq \exp(\lambda^2 \tau^2 / 2), \forall \lambda \in \mathbb{R}$.
- The control variates $h_1, \dots, h_m \in L^2(P)$ are **uniformly bounded**. Put $U_h := \max_{j=1, \dots, m} \sup_{x \in \mathcal{X}} |h_j(x)|$
- We have **orthogonality between active and inactive control variates** $P(h_j h_k) = 0$ for all $j \in \{1, \dots, m\} \setminus S^*$ and all $k \in S^*$.

Assumptions: **sub-gaussian** residuals, **linearly independent** and **bounded** control variates, appropriate λ

$$(U_h = \max_{j=1,\dots,m} \sup_{x \in \mathcal{X}} |h_j(x)|, \gamma = \lambda_{\min}(G))$$

Support recovery LASSOMC (Leluc, Portier, Segers, 2019)

For all $\delta \in (0, 1)$, all integer n such that $n \geq 70(\ell^* U_h^2 / \gamma^*)^2 \log(10\ell^* m / \delta)$, and all λ such that

$$17U_h \sqrt{\log(10m/\delta)\tau} / \sqrt{n} \leq \lambda \leq (\gamma^* / (3\sqrt{\ell^*})) \min_{k \in S^*} |\beta_k^*(f)|,$$

it holds that, with probability at least $1 - \delta$, the LASSO solution $\hat{\beta}_n^{\text{lasso}}(f)$ is unique and the true active set is recovered, $\text{supp}(\hat{\beta}_n^{\text{lasso}}(f)) = S^*$.